# On the Convergence Rate for a Kernel Estimate of the Regression Function

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## **Abstract**

We give the rate of the uniform convergence for the kernel estimate of the regression function over a sequence of compact sets which increases to  $\mathbb{R}^d$  when n approaches the infinity and when the observed process is  $\varphi$ -mixing. The used estimator for the regression function is the kernel estimator proposed by Nadaraya, Watson (1964).

**Keywords:** Kernel estimate,  $\varphi$ -Mixing, regression

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### 1. Introduction

Let  $(X_t, Y_t)_t \in \mathbb{N}$  be a strictly stationary process where  $(X_t, Y_t)$  takes on values in  $\mathbb{R}^d x \mathbb{R}$  and distributed as (X, Y). Suppose that a segment of data  $(X_t, Y_t)_{t=1}^n$  has been observed.

We are interested in the study of the rate of convergence for a kernel estimate of the regression function, known as:

$$r(x) = E(Y_t|X_t = x)$$
  $t \in \mathbb{N}$ .

A natural estimator of the function r(.) is given by

$$r_n(x) = \frac{\sum_{t=1}^{n} Y_t K\left(\frac{x - X_t}{h_n}\right)}{\sum_{t=1}^{n} K\left(\frac{x - X_t}{h_n}\right)}$$
  $\forall x \in E$ 

Where *E* stands for the subset  $\{x \in \mathbb{R}, f(x) > 0\}$ , *f* being the density of the process  $(X_t)$  and  $(h_n)$  is a positive sequence of real numbers such that  $h_n \to 0$  and  $nh_n^d \to \infty$  when  $n \to \infty$ .

K is a Parzen-Rosenblatt kernel type in the sense of a bounded function satisfying

$$\int_{\mathbb{R}} K(x)dx = 1 \quad \text{and} \quad \lim_{\|x\| \to \infty} \|x\| K(x) = 0$$

Moreover, it is assumed to be strictly positive and with bounded variation.

The estimation of the regression function has been subject to several investigations, and many authors have been involved. Among others, Devroye (1981), Collomb (1984, 1985), Györfy *et al.* (1989), Härdle (1990), Bosq (1996), Arfi (1996), Arfi (1997) and Walk (2006).

Watson (1964), for instance, considered the estimation of the conditional expectation as a predictor of Y and applied this method to some climatological time series data; Nadaraya (1964), established the same estimator independently.

Gasser *et al.* (1984) introduced a kernel estimate for obtaining a nonparametric estimate of a regression function and its derivatives, Sara Van De Ger (1990) proposed an entropy approach to establish rates of convergence for estimators of a regression function and later on Hermann and Ziegler (2004) studied the rates of consistency for a nonparametric estimation of the mode in absence of smoothness assumptions.

Our work is devoted to the rate of the uniform convergence for a kernel estimate of the regression function over an increasing sequence of compact sets under a mixing condition.

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# 2. Preliminaries and Assumptions

We assume that the process  $(X_t)_{t\in\mathbb{N}}$  is stationary and  $\varphi$ -mixing that is

$$\varphi_n = \sup_{A \in \mathcal{M}'_0} \sup_{B \in \mathcal{M}^{\infty}_{t+n}} \{ |P(B/A) - P(A)| \} \longrightarrow 0, \quad n \to \infty$$

where  $\mathcal{M}_0^t$  is the  $\sigma$ -field generated by  $\{X_0, X_1, ..., X_t\}$  and  $\mathcal{M}_{t+n}^{\infty}$  is the  $\sigma$ -field generated by  $\{X_{t+n}, X_{t+1+n}, ...\}$ 

We will make use of the following assumptions

A1. 
$$\exists \Gamma < \infty$$
,  $\forall x \in \mathbb{R}^d$ ,  $f(x) \leq \Gamma$ 

and

$$\exists \gamma_n > 0, \quad \forall x \in C_n, \quad f(x) \ge \gamma_n$$

where  $C_n$  is a sequence of compact sets such that  $C_n = \{x : ||x|| \le c_n\}$  and  $c_n \to \infty$ 

A2.  $\exists \beta \geq 2$ ,  $\exists M < \infty$  such that  $E(|Y|^{\beta}) \leq M$ 

A3. 
$$\exists V < \infty$$
,  $\forall x \in \mathbb{R}^d$ ,  $E[(Y - r(x))^2 | X = x] \le V$ 

A4. The density f is twice differentiable and its second derivatives are bounded on  $\mathbb{R}^d$ 

A5. The kernel *K* is Lipschitz of ratio  $L_k$  that is  $|K(x) - K(y)| \le L_k ||x - y||^{\gamma_1}$ 

### 3. Main Result

#### Theorem

Assuming that the assumptions A1 through A5 hold, we further assume that the function r is Lipschitz, bounded on  $\mathbb{R}^d$  and that the bandwith sequence  $(h_n)$  satisfies with  $y_n$ :

$$n^{\delta} \gamma_n^{-1} h_n^{-d} y_n^{-\beta} \longrightarrow \infty \quad n \to \infty$$

$$\forall \epsilon_0 > 0 \quad \sum_n \left\{ \frac{c_n^d y_n^{d/\gamma_1} (log n)^{d/\gamma_1}}{h_n^{d(1+d/\gamma_1)}} \right\} \exp\left(-\epsilon_0^2 \frac{n^{1-2\delta} y_n^2 h_n^d}{m_n y_n}\right) < \infty$$

Where  $\delta \in ]0, 1/2[$ ,  $m_n$  and  $y_n$  are two sequences such that:

$$1 \le m_n \le n/2$$
 and  $1 \le y_n \le \sqrt{n/2}$ 

If the kernel K is even with  $\int z^2 K(z) dz < \infty$  for  $z = (z_1, ..., z_d)$  and if there exists a constant D such that  $\gamma_n^{-1} y_n n^{\delta} h_n^d < D$  then if r is continuous, Lipschitz and bounded on  $\mathbb{R}$  we have:

$$n^{\delta} \sup_{\|x\| \le c_n} |r_n(x) - r(x)| = O(1) \quad a.s. \quad n \to \infty.$$

# 4. Preliminary Results

For practical reason, we make the following decomposition:

$$r_n(x) - r(x) = \frac{1}{f(x)} \{ [g_n(x) - r(x)f(x)] - r_n(x) [f_n(x) - f(x)] \}$$

where 
$$g_n(x) = \frac{1}{nh_n^d} \sum_{t=1}^n Y_t K\left(\frac{x - X_t}{h_n}\right)$$

and 
$$f_n(x) = \frac{1}{nh_n^d} \sum_{t=1}^n K\left(\frac{x - X_t}{h_n}\right)$$

This leads to

$$\sup_{x \in C_n} |r_n(x) - r(x)| = \frac{1}{f(x)} \left\{ \sup_{x \in C_n} |g_n(x) - r(x)f(x)| + \sup_{x \in C_n} |r_n(x)| |f_n(x) - f(x)| \right\}$$

Then if

$$\sup_{x \in C_n} |r_n(x)| \le y_n \quad a.s. \quad \text{we obtain}$$

$$\sup_{x \in C_n} |r_n(x) - r(x)| = \gamma_n^{-1} \left\{ \sup_{x \in C_n} |g_n(x) - r(x)f(x)| + y_n \sup_{x \in C_n} |f_n(x) - f(x)| \right\}$$

The following Lemma will be used in our proofs

Lemma (Collomb, 1984)

Let  $Z_t$  be a real centered and  $\varphi$ -mixing process such tha:

$$|Z_t| \le d_n$$
  $E(Z_t^2) \le D_n$   $E|Z_t| \le \delta_n$ 

then,  $\forall \epsilon > 0$ ,  $\forall n \in \mathbb{N}^*$  we have:

$$P\left\{\left|\sum_{t=1}^{n} Z_{t}\right| > \epsilon\right\} 2 \exp\left\{-\alpha\epsilon + 3\sqrt{e}n\frac{\varphi_{m}}{m} + 6\alpha^{2}n\left[D_{n} + 6\delta_{n}d_{n}\sum_{j=1}^{m}\varphi_{j}\right]\right\}$$

# Lemma 1

Under the hypotheses of Theorem we have:

$$\gamma_n^{-1} n^{\delta} \sup_{x \in C_n} |g_n(x) - Eg_n(x)| \to 0 \quad a.s. \quad n \to \infty.$$

Proof:

Because of the possible large values for  $Y_t$ , we use a truncation technique which consists in decomposing  $g_n$  in  $g_n^+$  and  $g_n^-$  where

$$g_n^+(x) = \frac{1}{nh_n^d} \sum_{t=1}^n Y_t \mathbb{I}_{[|Y_t| > y_n]} K\left(\frac{x - X_t}{h_n}\right)$$

and  $g_n^-(x) = g_n(x) - g_n^+(x)$ , where  $y_n$  is the unbounded sequence defined in the Theorem.

We start by showing that:

$$\gamma_n^{-1} n^{\delta} \sup_{\|x\| \le c_n} |g_n^-(x) - Eg_n^-(x)| \to 0 \quad a.s. \quad n \to \infty.$$

To this end, we write:

$$\begin{split} g_n^-(x) - Eg_n^-(x) &= \sum_{t=1}^n \varphi_t \quad \text{with} \\ \varphi_t &= \frac{1}{nh_n^d} \left\{ Y_t \mathbb{I}_{[|Y_t| \leq y_n]} K\left(\frac{x - X_t}{h_n}\right) - E\left[ Y_t \mathbb{I}_{[|Y_t| \leq y_n]} K\left(\frac{x - X_t}{h_n}\right) \right] \right\} \end{split}$$

therefore  $E(\varphi_t) = 0$ ;

 $|\varphi_t| \le \frac{2K_1 y_n}{nh_n^d} = d_n$  where  $K_1$  is an upperbound of K, which permits to write:

$$E|\varphi_t| \le \frac{2\Gamma}{n} E\left| \frac{Y_t}{h_n^d} \mathbb{I}_{[|Y_t| \le y_n]} K(\frac{x - X_t}{h_n}) \right| \le \frac{2\Gamma}{n} \int \frac{E(|Y_t|/X_t = u)}{h_n^d} K\left(\frac{x - u}{h_n}\right) du$$

Leading by Schwartz inequality and the assumption A3 to:

$$|E|\varphi_t| \le \frac{2\Gamma}{n} \int \frac{(r^2(u) + V)^{1/2}}{h_n^d} K\left(\frac{x - u}{h_n}\right) du \le \tau_1 n^{-1}$$

where  $\tau_1$  is a positive constant.

Now, same arguments give:

$$E(\varphi_t)^2 \le \frac{2\Gamma}{n^2} \int \frac{(r^2(u) + V)}{h_n^{2d}} K\left(\frac{x - u}{h_n}\right) du \le \upsilon n^{-2} h_n^{-d}$$

where v is a positive constant.

We apply the Collomb inequality with  $\alpha = 1/(4m_nd_n)$  and we obtain for all  $\epsilon_n > 0$ :

$$P(|g_n^-(x) - Eg_n^-(x)| > \epsilon_n) \le C_1 \exp\left(-C_2\epsilon_n^2 \frac{nh_n^d}{m_n y_n}\right)$$

where  $C_1$  and  $C_2$  are two positive constants.

Next, we cover  $C_n$  by  $\mu_n^d$  spheres in the shape of  $\{x: ||x-x_{jn}|| \le c_n \mu_n^{-1}\}$  where  $1 \le j \le \mu_n^d$ 

And we make the following decomposition:

$$|g_n^-(x) - Eg_n^-(x)| \le |g_n^-(x) - g_n^-(x_{jn})| + |g_n^-(x_{jn}) - Eg_n^-(x_{jn})| + |Eg_n^-(x_{jn}) - g_n^-(x)|$$

and we have

$$|g_n^-(x) - g_n^-(x_{jn})| \le \frac{y_n}{nh_n^d} \sum_{t=1}^n \left| K\left(\frac{x - X_t}{h_n}\right) - K\left(\frac{x_{jn} - X_t}{h_n}\right) \right|$$

The kernel *K* being Lipschitz we obtain

$$|g_n^-(x) - g_n^-(x_{jn})| \le L_K \frac{y_n}{h_n^{d+\gamma_1}} ||x - x_{jn}||^{\gamma_1}$$

$$|g_n^-(x) - g_n^-(x_{jn})| \le L_K \frac{y_n}{h_n^{d+\gamma_1}} c_n^{\gamma_1} \mu_n^{-\gamma_1}$$

$$|g_n^-(x) - g_n^-(x_{jn})| \le \frac{1}{\log n}$$

If we choose

$$\mu_n = L_K^{1/\gamma_1} \frac{y_n^{1/\gamma_1} c_n (log n)^{1/\gamma_1}}{h_n^{d/\gamma_1+1}} \longrightarrow \infty.$$

Thus we obtain:

$$\sup_{x \in C_n} |g_n^-(x) - Eg_n^-(x)| \le \sup_{1 \le j \le \mu_n^d} |g_n^-(x_{jn}) - Eg_n^-(x_{jn})| + \frac{2}{\log n}$$

Therefore, if we apply  $\mu_n^d$  times the Lemma of Collomb, we obtain

$$P\left(\sup_{x\in C_n}|g_n^-(x)-Eg_n^-(x)|>2\epsilon_n\right)\leq C_1\mu_n^d\exp\left(-C_2\epsilon_n^2\frac{nh_n^d}{m_ny_n}\right)$$

Now if we choose  $\epsilon_n = n^{-\delta} \gamma_n \epsilon_0$  for a certain  $\epsilon_0 > 0$ , we obtain accordingly with the hypotheses of the Theorem:

$$P\left(\gamma_n^{-1} n^{\delta} \sup_{\|x\| \le c_n} |g_n^-(x) - Eg_n^-(x)| > 2\epsilon_0\right) \le C_1 L_K^{d/\gamma_1} \frac{y_n^{d/\gamma_1} c_n^d (logn)^{d/\gamma_1}}{h_n^{d(1+d/\gamma_1)}} \exp\left(-C_2 \epsilon_0^2 \frac{n^{1-2\delta} \gamma_n^2 h_n^d}{m_n y_n}\right)$$

The hypotheses of the Theorem permit to conclude that:

$$\gamma_n^{-1} n^{\delta} \sup_{\|\mathbf{x}\| \le c_n} |g_n^-(\mathbf{x}) - Eg_n^-(\mathbf{x})| \longrightarrow 0, \quad a.s. \quad n \to \infty$$

It remains to show that:

$$n^{\delta} \gamma_n^{-1} \sup_{\|x\| \le c_n} |g_n^+(x) - Eg_n^+(x)| \to 0, \quad a.s. \quad n \to \infty.$$

For practical reason, we write:

$$n^{\delta} \gamma_n^{-1} \sup_{\|x\| \le c_n} |g_n^+(x) - Eg_n^+(x)| \le E_n + F_n$$

Where,

$$E_n = \frac{n^{\delta} \gamma_n^{-1}}{n h_n^d} \sup_{\|x\| \le c_n} \left| \sum_{t=1}^n Y_t \mathbb{I}_{(|Y_t| > y_n)} K\left(\frac{x - X_t}{h_n}\right) \right|$$

And we have

$$(E_n \neq 0) \subset \{\exists t_0 \in [1, 2, 3, ..., n] \text{ such that } |Y_{t_0}| > y_n\}$$

the above leads to

$$(E_n \neq 0) \subset \bigcup_{t=1}^n \{|Y_t| > y_n\}$$

$$P(E_n \neq 0) \le \sum_{t=1}^{n} P(|Y_t| > y_n) = nP(|Y_t| > y_n)$$

$$\sum_{t=1}^{n} P(E_t \neq 0) \le \sum_{t=1}^{n} P(|Y_t| > y_n) \le \sum_{t=1}^{n} ny_n^{-\beta} E|Y|^{\beta}$$

$$\sum_{t=1}^{n} P(E_t \neq 0) \le c_4 \sum_{t=1}^{n} ny_n^{-\beta} < \infty$$

where  $c_4$  is a positive constant.

Then  $E_n \to 0$ , a.s.,  $n \to \infty$  and  $\sup_{1 \le t \le n} |Y_t| \le y_n$  a.s.

The kernel K being strictly positive, we conclude that  $|r_n(x)| \le y_n$  a.s.

Moreover,

$$\begin{split} F_n &= \frac{n^{\delta}}{\gamma_n n h_n^d} \sup_{\|x\| \leq c_n} \left| \sum_{t=1}^n E\left[ Y_t \mathbb{I}_{[|Y_t| > y_n]} K\left(\frac{x - X_t}{h_n}\right) \right] \right| \\ & F_n \leq \frac{n^{\delta}}{\gamma_n h_n^d} K_1 E[|Y| \mathbb{I}_{|Y| > y_n}] \\ F_n &\leq \frac{n^{\delta}}{\gamma_n h_n^d} K_1 (E(Y^2))^{1/2} (P[|Y| > y_n])^{1/2} \leq c_5 n^{\delta} \gamma_n^{-1} h_n^{-d} y_n^{-\beta/2} \to 0, \quad n \to \infty \end{split}$$

where  $c_5$  is being a positive constant.

### Lemma 2

Under the assumptions of the Theorem, we have:

$$n^{\delta} \gamma_n^{-1} \sup_{x \in \mathbb{R}^d} |Eg_n(x) - r(x)f(x)| \to 0, \quad n \to \infty.$$

Proof:

$$Eg_n(x) - r(x)f(x) = \frac{1}{nh_n^d} E\left\{\sum_{t=1}^n Y_t K\left(\frac{x - X_t}{h_n}\right)\right\} - r(x)f(x)$$

$$Eg_n(x) - r(x)f(x) = \frac{1}{h_n^d} \int_{\mathbb{R}^d} r(u)K\left(\frac{x - u}{h_n}\right)f(u)du - r(x)f(x)$$

We write  $z = (x - u)/h_n$  and we obtain

$$Eg_{n}(x) - r(x)f(x) = \int_{\mathbb{R}^{d}} [r(x - zh_{n}) - r(x)]K(z)f(x - zh_{n})dz + r(x) \int K(z)[f(x - zh_{n}) - f(x)]dz$$

Assuming that the function r(.) is Lipschitz of ratio 1 and order 1 provides

$$\left| \int_{\mathbb{R}^d} [r(x - zh_n) - r(x)] K(z) f(x - zh_n) dz \right| \le h_n \Gamma \int |Z| K(z) dz$$

Now a Taylor expansion, the Bochner lemma and the fact that the function r is bounded permit to conclude that

$$n^{\delta} \gamma_n^{-1} \sup_{\mathbf{y} \in \mathbb{R}^d} |Eg_n(\mathbf{x}) - r(\mathbf{x})f(\mathbf{x})| \to 0, \quad n \to \infty.$$

# Lemma 3

Under the assumptions of the Theorem, we have:

$$\lim_{n\to\infty} \frac{y_n n^{\delta}}{\gamma_n} \sup_{\|x\| \le c} |f_n(x) - Ef_n(x)| = 0 \quad a.s.$$

Proof:

This is a particular case of Lemma 1 when  $Y_t = 1$  and  $\epsilon = \epsilon_0 \gamma_n n^{-\delta} y_n^{-1}$  for a certain  $\epsilon_0 > 0$ .

### Lemma 4

Under the assumptions of the Theorem, we have:

$$\lim_{n \to \infty} \frac{y_n n^{\delta}}{\gamma_n} \sup_{x \in \mathbb{R}^d} |Ef_n(x) - f(x)| = 0$$

Proof:

We write

$$Ef_n(x) - f(x) = \frac{1}{h_n^d} \int [f(u) - f(x)] K\left(\frac{u - x}{h_n}\right) du$$

A Taylor expansion and the hypotheses of the Theorem and the Bochner lemma permit to conclude.

# 5. Proof of the Theorem

The lemma 1, Lemma 2, Lemma 3, and Lemma 4 permit to conclude.

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#### References

- Arfi, M. (1996). Sur la Regression non parametrique d'un processus stationnaire mélangeant ou ergodique. Doctorate Thesis of University Paris 6.
- Arfi, M. (1997). On General Estimate of the Regression Function. *Statistics & Decisions*, 15, 191-200. http://dx.doi.org/10.1524/strm.1997.15.2.191
- Bosq, D. (1996). Nonparametric statistics for stochastic processes. *Lecture Notes in Statistics*, 110, Springer Verlag. http://dx.doi.org/10.1007/978-1-4684-0489-0
- Collomb, (1984). Proprietes de Convergence Presque Sure du Predicteur a Noyau. Z. Wahrsch. Verw., Gebiete., 66, 441-460. http://dx.doi.org/10.1007/BF00533708
- Collomb, G. (1985) Nonparametric regression; an up to date bibliography. *Statistics*, *16*, 309-324. http://dx.doi.org/10.1080/02331888508801860
- Devroye, L. (1981). On the Almost Eveywhere Convergence of Nonparametric Regression Function Estimates. *The Annals of Statistics*, 9, 1310-1319. http://dx.doi.org/10.1214/aos/1176345647
- Gasser, T., & Hans-Georg Müller. (1984). Estimating Regression Functions and Their Derivatives by the Kernel Method. *Scand. J. Statist.*, 11, 171-185.
- Györfy, L., Härdle, W., Sarda, P., & Vieu, P. (1989). Nonparametric Curve Estimation from Time Series. *Lecture Notes in Statistics*, 60, Springer Verlag.
- Härdle, W. (1990). Applied Nonparametric Regression. Cambridge Univ. Press.
- Hermann, E., & Ziegler, K. (2004). Rates of Consistency for a Nonparametric Estimation of the Mode in Absence of Smoothness Assumptions. *Statistics & Probability Letters*, 68, 359-368. http://dx.doi.org/10.1016/j.spl.2004.04.005
- Nadaraya, E. A. (1964). On estimating regression. *Theory of Probability and its Application*, 9, 141-142. http://dx.doi.org/10.1137/1109020
- Sara Van De Ger. (1990). Estimating a Regression Function. *The Annals of Statistics*, 18(2), 907-924. http://dx.doi.org/10.1214/aos/1176347632
- Walk, H. (2006). Almost sure Convergence Properties of Nadaray-Watson Regression Estimates. *International Series in Operations Research & Management Science*, 46, Modeling Uncertainty, 201-223.
- Watson, G. S. (1964) Smooth regression analysis. Sankhya, A 26, 359-372.

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