# Olsavs: A New Algorithm For Model Selection 

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#### Abstract

The shrinkage methods such as Lasso and Relaxed Lasso introduce some bias in order to reduce the variance of the regression coefficients in multiple linear regression models. One way to reduce bias after shrinkage of the coefficients would be to apply ordinary least squares to the subset of predictors selected by the shrinkage method used. This work extensively investigated this idea and developed a new variable selection algorithm. The authors named this technique OLSAVS (Ordinary Least Squares After Variable Selection). The OLSAVS algorithm was implemented in $R$. Simulations were used to illustrate that the new method is able to produce better predictions with less bias for various error distributions. The OLSAVS method was compared with a few widely used shrinkage methods in terms of their achieved test root mean square error and bias.


Keywords: Multiple Linear Regression, OLS, Lasso, Relax Lasso, Elastic Net, Bias, Variance

## 1. Introduction

Following (Pelawa Watagoda et al., 2021) and (Pelawa Watagoda, 2018), suppose that the response variable $Y_{i}$ and at least one predictor variable $x_{i, j}$ are quantitative with $x_{i, 1} \equiv 1$. Let $\boldsymbol{x}_{i}^{T}=\left(x_{i, 1}, \ldots, x_{i, p}\right)=\left(\begin{array}{ll}1 & \boldsymbol{u}_{i}^{T}\end{array}\right)$ and $\boldsymbol{\beta}=\left(\beta_{1}, \ldots, \beta_{p}\right)^{T}$ where $\beta_{1}$ corresponds to the intercept. Then the multiple linear regression (MLR) model is

$$
\begin{equation*}
Y_{i}=\beta_{1}+x_{i, 2} \beta_{2}+\cdots+x_{i, p} \beta_{p}+e_{i}=\boldsymbol{x}_{i}^{T} \boldsymbol{\beta}+e_{i} \tag{1}
\end{equation*}
$$

for $i=1, \ldots, n$. This model is also called the full model. Here $n$ is the sample size, and assume that the random variables $e_{i}$ are independent and identically distributed (iid) with variance $V\left(e_{i}\right)=\sigma^{2}$.
In matrix notation, these $n$ equations become

$$
\begin{equation*}
Y=X \beta+e \tag{2}
\end{equation*}
$$

where $\boldsymbol{Y}$ is an $n \times 1$ vector of response variables, $\boldsymbol{X}$ is an $n \times p$ matrix of predictors, $\boldsymbol{\beta}$ is a $p \times 1$ vector of unknown coefficients, and $\boldsymbol{e}$ is an $n \times 1$ vector of unknown errors.

$$
\left[\begin{array}{c}
y_{1}  \tag{3}\\
y_{2} \\
\vdots \\
y_{n}
\end{array}\right]=\left[\begin{array}{ccccc}
1 & x_{12} & x_{13} & \ldots & x_{1 p} \\
1 & x_{22} & x_{23} & \ldots & x_{2 p} \\
\vdots & \vdots & \vdots & \ddots & \vdots \\
1 & x_{n 2} & x_{n 3} & \ldots & x_{n p}
\end{array}\right] \times\left[\begin{array}{c}
\beta_{1} \\
\beta_{2} \\
\vdots \\
\beta_{p}
\end{array}\right]+\left[\begin{array}{c}
e_{1} \\
e_{2} \\
\vdots \\
e_{n}
\end{array}\right]
$$

The $i$ th fitted value $\hat{Y}_{i}=\boldsymbol{x}_{i}^{T} \hat{\boldsymbol{\beta}}$ and the $i$ th residual $r_{i}=Y_{i}-\hat{Y}_{i}$ where $\hat{\boldsymbol{\beta}}$ is an estimator of $\boldsymbol{\beta}$.

## 2. Variable Selection

Variable selection is the search for a subset of predictor variables that can be deleted with little loss of information if $n / p$ is large, and so the model with the remaining predictors is useful for prediction. Following (Olive and Hawkins, 2005) and (Pelawa Watagoda and Olive, 2021), a model for variable selection can be described by

$$
\begin{equation*}
\boldsymbol{x}^{T} \boldsymbol{\beta}=\boldsymbol{x}_{S}^{T} \boldsymbol{\beta}_{S}+\boldsymbol{x}_{E}^{T} \boldsymbol{\beta}_{E}=\boldsymbol{x}_{S}^{T} \boldsymbol{\beta}_{S} \tag{4}
\end{equation*}
$$

where $\boldsymbol{x}=\left(\boldsymbol{x}_{S}^{T}, \boldsymbol{x}_{E}^{T}\right)^{T}, \boldsymbol{x}_{S}$ is an $a_{S} \times 1$ vector, and $\boldsymbol{x}_{E}$ is a $\left(p-a_{S}\right) \times 1$ vector. Given that $\boldsymbol{x}_{S}$ is in the model, $\boldsymbol{\beta}_{E}=\mathbf{0}$ and $E$ denotes the subset of terms that can be eliminated given that the subset $S$ is in the model. Let $\boldsymbol{x}_{I}$ be the vector of $a$ terms from a candidate subset indexed by $I$, and let $\boldsymbol{x}_{O}$ be the vector of the remaining predictors (out of the candidate
submodel). Suppose that $S$ is a subset of $I$ and that model (5) holds. Then

$$
\begin{equation*}
\boldsymbol{x}^{T} \boldsymbol{\beta}=\boldsymbol{x}_{S}^{T} \boldsymbol{\beta}_{S}=\boldsymbol{x}_{S}^{T} \boldsymbol{\beta}_{S}+\boldsymbol{x}_{I / S}^{T} \boldsymbol{\beta}_{(I / S)}+\boldsymbol{x}_{O}^{T} \mathbf{0}=\boldsymbol{x}_{I}^{T} \boldsymbol{\beta}_{I} \tag{5}
\end{equation*}
$$

where $\boldsymbol{x}_{I / S}$ denotes the predictors in $I$ that are not in $S$. Since this is true regardless of the values of the predictors, $\boldsymbol{\beta}_{O}=\mathbf{0}$ if $S \subseteq I$.

## 3. Estimating Model Coefficients

The most common method of obtaining model coefficients $(\boldsymbol{\beta})$ is the ordinary least squares. There are many methods for estimating $\boldsymbol{\beta}$, including, Lasso by (Tibshirani, 1996), Elastic Net by (Zou and Hastie, 2005), Relaxed Lasso by (Meinshausen, 2007), and ridge regression by (Hoerl and Kennard, 1970).
One can obtain the the least squares estimates for $\beta_{1}, \beta_{1}, \ldots, \beta_{p}$ by minimizing (6)

$$
\begin{equation*}
Q=\sum_{i=1}^{n}\left(Y_{i}-\beta_{1}-\beta_{2} X_{i, 2}-\ldots-\beta_{p} X_{i, p}\right)^{2} \tag{6}
\end{equation*}
$$

Ridge Regression coefficient estimates $\hat{\beta}^{R}$, are values that minimizes,

$$
\begin{equation*}
\sum_{i=1}^{n}\left(y_{i}-\beta_{1}-\sum_{j=2}^{p} \beta_{j} X_{i j}\right)^{2}+\lambda \sum_{j=2}^{p} \beta_{j}^{2}=R S S+\lambda \sum_{j=2}^{p} \beta_{j}^{2} \tag{7}
\end{equation*}
$$

Where $\lambda \geq 0$ is a tuning parameter. The term, $\lambda \sum_{j=2}^{p} \beta_{j}^{2}$ is known as the shrinkage penalty. This penalty value is small when $\beta_{1}, \ldots, \beta_{p}$ are close to zero, sending the $\beta_{j}$ values to zero but never reaching zero. For this reason, ridge regression includes all predictors $p$ in the model. The Lasso regression minimizes a similar quantity as in (7), except the shrinkage penalty changed to $\lambda \sum_{j=1}^{p}\left|\beta_{j}\right|$. Unlike ridge regression, often, some of the Lasso coefficients $\hat{\beta_{j}}$ are exactly equal to zero. Following (Meinshausen, 2007), Relaxed Lasso controls model selection and shrinkage estimation by two separate parameters $\lambda$ and $\phi$. The Relaxed Lasso estimator is defined for $\lambda \in[0, \infty)$ and $\phi \in(0,1]$ as

$$
\begin{equation*}
\hat{\boldsymbol{\beta}}^{\lambda, \phi}=\underset{\beta}{\operatorname{argmin}} n^{-1} \sum_{i=1}^{n}\left(\boldsymbol{Y}_{i}-\boldsymbol{X}_{i}^{T}\left\{\boldsymbol{\beta} \cdot \mathbf{1}_{\mathcal{M}_{\lambda}}\right\}\right)^{2}+\phi \lambda|\boldsymbol{\beta}|_{1} \tag{8}
\end{equation*}
$$

Where $\mathbf{1}_{\mathcal{M}_{\lambda}}$ is the indicator function on the set of variables $\mathcal{M}_{\lambda} \subseteq\{1, \ldots, p\}$ so that for all $k \in\{1, \ldots, p\}$

$$
\boldsymbol{\beta} \cdot \mathbf{1}_{\mathcal{M}_{\lambda}}= \begin{cases}0 & k \notin \mathcal{M}_{\lambda} \\ \boldsymbol{\beta}_{k} & k \in \mathcal{M}_{\lambda}\end{cases}
$$

The Elastic Net estimator is defined as follows:
Given dataset $(\boldsymbol{y}, \boldsymbol{X})$, penalty parameter $\left(\lambda_{1}, \lambda_{2}\right)$ and augmented data $\left(\boldsymbol{y}^{*}, \boldsymbol{X}^{*}\right)$

$$
\boldsymbol{X}_{(n+p) \times p}^{*}=\left(1+\lambda_{2}\right)^{(-1 / 2)}\binom{\boldsymbol{X}}{\sqrt{\lambda_{2}} \boldsymbol{I}}, \quad \boldsymbol{y}_{(n+p)}^{*}=\binom{\boldsymbol{y}}{\mathbf{0}}
$$

Let $\gamma=\frac{\lambda_{1}}{\sqrt{1+\lambda_{2}}}$ and $\boldsymbol{\beta}^{*}=\sqrt{1+\lambda_{2}} \boldsymbol{\beta}$. Then the naïve Elastic Net solves a Lasso-type problem

$$
\begin{equation*}
\hat{\boldsymbol{\beta}}^{*}=\underset{\boldsymbol{\beta}^{*}}{\operatorname{argmin}} n^{-1} \sum_{i=1}^{n}\left(\boldsymbol{Y}_{i}^{*}-\boldsymbol{X}_{i}^{* T} \boldsymbol{\beta}^{*}\right)^{2}+\frac{\lambda_{1}}{\sqrt{1+\lambda_{2}}}\left|\boldsymbol{\beta}^{*}\right|_{1} \tag{9}
\end{equation*}
$$

Naïve Elastic Net estimator is a two steps procedure: for each fixed $\lambda_{2}$ it first finds the ridge regression coefficients, and then it does the Lasso type shrinkage along the Lasso type solution path. As a result, the predictors will shrink unnecessarily (double shrinkage). This would not help to reduce the variances much and also it will introduce unnecessary extra bias, compared to the original Lasso or ridge. As a solution, it uses a correction factor $\sqrt{1+\lambda_{2}}$ to get the Elastic Net solutions.

Finally, the Elastic Net solutions can be written as $\hat{\boldsymbol{\beta}}=\sqrt{1+\lambda_{2}} \hat{\boldsymbol{\beta}}^{*}$.

## 4. A New Method for Model Selection: OLSAVS

The shrinkage methods such as Lasso and Relaxed Lasso introduce some bias in order to reduce the variance of the regression coefficients. As briefly mentioned in (Hastie et al., 2015), one way to reduce bias after shrinkage of the coefficients would be to apply ordinary least squares to the subset of predictors selected by the shrinkage method used. This work extensively explores this idea to develop a new variable selection algorithm. The authors named this technique Ordinary Least Squares After Variable Selection (OLSAVS). OLSAVS method was implemented in $R$. The set of functions can be found at https://hasthika.github.io/olsvspack.txt. The algorithm of the OLSAVS method is as follows:

```
Algorithm: Ordinary Least Squares After Variable Selection (OLSAVS)
    Repeat: following steps with a different shrinkage method
    1) Apply the first shrinkage method to \(\left(Y_{i}, \boldsymbol{x}_{i}\right)\) for \(i=1, \ldots, n\).
    2) Obtain the \(k\) non-zero predictors selected by the shrinkage method in 1)
    3) Apply Ordinary Least Squares on the subset of \(k\) predictors obtained in 2)
```


## Stop

4) Select a single best model using cross-validated prediction error, $C_{p}$, (AIC), BIC, or adjusted $R^{2}$

## 5. Simulation

This section contains the simulation setup and the results.

### 5.1 Simulation Setup

The statistical software, R (see ( R Core Team, 2020)) was used to generate $\left(Y_{i}, \boldsymbol{x}_{i}\right)$ for $i=1, \ldots, n$. The regression parameters $\beta$ were set to $(1,1, \ldots, 1,0, \ldots, 0)$ with $k+1$ ones, $p-k-1$ zeroes where $p$ is the total number of predictors, and $k$ is the number of non-trivial predictors. Then for a given regression method, the regression coefficients, $\hat{\boldsymbol{\beta}}$ were obtained using the proposed method. This process was repeated 5000 times (runs). For each run, the difference between the regression parameters $\boldsymbol{\beta}$ and the regression coefficients, $\hat{\boldsymbol{\beta}}$ were obtained using the Minkowski distance. The average difference (Diff) was calculated by averaging all 5000 runs. The test root mean square error was also obtained using a set of test observations and averaged over the 5000 runs (TRMSE). $p=n / 5, n / 2$, or $n-1$ were used as the total number of predictors and $k=1,19$, or $p-1$ as the number of non-trivial predictors in the model. As per the easiness of coding the relation $\operatorname{cor}\left(x_{i}, x_{j}\right)=\rho=\left(2 \psi+(p-3) \psi^{2}\right) /\left(1+(p-2) \psi^{2}\right)$ was used, for $i \neq j$, where, $x_{i}, x_{j}$ are non-trivial predictors. As $\psi$ increases the correlation between preceptors, $\rho$ grows. $\psi=0,0.3$ or 0.9 were used with five error distributions with zero mean.

1. $N(0,1)$, the normal distribution with mean 0 and variance 1 which is commonly used in simulation studies.
2. $t_{3}$, a $t$ distribution with degrees of freedom 3 , one of the heavy-tailed distributions.
3. $E X P(1)-1$, an exponential distribution with mean 0 . This distribution is not very commonly used in simulations but found in many real-life situations, a non-symmetric error distribution
4. uniform $(-1,1)$, a uniform distribution in the rage of -1 and 1 .
5. $0.9 N(0,1)+0.1 N(0,100)$, a mixture of normal distributions.

The simulation study was conducted in $R$.

### 5.2 Simulation Results

Table 1 compares the OLSAVS method to Lasso with normal errors with mean 0 and variance 1 . Notice in the TRMSE column, when the number of non-trivial predictors $(k)$ is low, Lasso and the OLSAVS method perform equally well. However, as $k$ and the correlation between the predictors increase, the OLSAVS method outperformed Lasso with noticeably larger distances between the OLSAVS and Lasso TRMSE values. This trend continues throughout the table. The OLSAVS stays consistent throughout for the TRMSE values, whereas the Lasso shows a lot of variability. Except for the case of $k=1$ and $\psi=0.9$, the Lasso either came close or betters OLSAVS in the difference value (Diff column). Other than this certain case, the OLSAVS method significantly bettered the Lasso.
The side-by-side boxplots in Figure 1, compare the TRMSE results for two randomly selected rows in table 1 for $n=100$ and $n=200$. OLSAVS has a smaller median TRMSE and a lower variation in the results than the Lasso. It also appears

Table 1. TRMSE and difference values for OLSAVS vs. Lasso for $e_{i} \sim N(0,1)$

|  |  |  |  | TRMSE |  | Diff |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| n | p | k | $\psi$ | OLSAVS | Lasso | OLSAVS | Lasso |
| 100 | 20 | 1 | 0 | 1.0478 | 1.021 | 0.01657 | 0.1449 |
| 100 | 20 | 1 | 0.3 | 1.0417 | 1.0199 | 0.1124 | 0.1549 |
| 100 | 20 | 1 | 0.9 | 1.0083 | 1.0057 | 0.6455 | 0.6635 |
| 100 | 20 | 19 | 0 | 1.1112 | 1.1107 | 0.0037 | 0.0122 |
| 100 | 20 | 19 | 0.3 | 1.1112 | 1.1377 | 0.0055 | 0.0158 |
| 100 | 20 | 19 | 0.9 | 1.1543 | 2.3451 | 1.5667 | 7.5674 |
| 100 | 50 | 1 | 0 | 1.0916 | 1.0369 | 0.0268 | 0.181 |
| 100 | 50 | 1 | 0.3 | 1.0708 | 1.0303 | 0.1654 | 0.2028 |
| 100 | 50 | 1 | 0.9 | 1.0142 | 1.0059 | 0.9511 | 0.5929 |
| 100 | 50 | 49 | 0 | 1.3973 | 1.4014 | 0.0043 | 0.0149 |
| 100 | 50 | 49 | 0.3 | 1.5522 | 4.353 | 0.0163 | 0.0557 |
| 100 | 50 | 49 | 0.9 | 2.1118 | 9.4775 | 8.8016 | 43.9716 |
| 200 | 40 | 1 | 0 | 1.0386 | 1.0175 | 0.0225 | 0.1208 |
| 200 | 40 | 1 | 0.3 | 1.0316 | 1.0157 | 0.1044 | 0.1327 |
| 200 | 40 | 1 | 0.9 | 1.0108 | 1.005 | 0.7825 | 0.4165 |
| 200 | 40 | 39 | 0 | 1.117 | 1.119 | 0.002 | 0.0095 |
| 200 | 40 | 39 | 0.3 | 1.117 | 2.5606 | 0.0028 | 0.0369 |
| 200 | 40 | 39 | 0.9 | 1.6985 | 6.8533 | 5.8545 | 32.8388 |
| 200 | 100 | 1 | 0 | 1.0691 | 1.0261 | 0.0354 | 0.1454 |
| 200 | 100 | 1 | 0.3 | 1.0605 | 1.0251 | 0.1392 | 0.1653 |
| 200 | 100 | 1 | 0.9 | 1.0184 | 1.0045 | 0.9712 | 0.0855 |
| 200 | 100 | 99 | 0 | 1.4495 | 1.4561 | 0.0034 | 0.0132 |
| 200 | 100 | 99 | 0.3 | 3.166 | 11.4553 | 0.1379 | 0.488 |
| 200 | 100 | 99 | 0.9 | 3.7274 | 27.9646 | 17.4699 | 93.9076 |



Figure 1. Box plots for table 1 showing simulation results for OLSAVS vs. Lasso regression with error type 1
for this simulation, for Lasso with normal errors, when the sample size increases, so does the variation in the regression coefficients while OLSAVS maintains the same low variability.
Table 2 has the same structure as in Table 1 but the simulation was done with error type 2. The Lasso seems to narrowly outperform OLSAVS when $k=1$. However, as the $k$ increases, the TRMSE favored the OLSAVS method significantly. Additionally, as the correlation between the non-trivial predictors increases, notice the large distance increase between the two methods in the TRMSE. In the Diff column in Table 2, the OLSAVS method outperformed Lasso by a large majority.
Figure 2 shows two simulation plots pulled from table 2. Looking at two cases, each varying in sample size, the OLSAVS does edge-out Lasso regression with the error type being from a t-distribution. However, unlike in Figure 1, the variation in the plot now decreases as expected when the sample size increases.
Table 3 compares OLSAVS with Lasso with the error being from the exponential distribution. A similar trend occurs in the TRMSE column as in Table 2. As the non-trivial predictors are increased to $p-1$ or the correlation between the predictors increased, OLSAVS produced lower TRMSE than Lasso. The OLSAVS seems to dominate the majority of the difference values.

Figure 3 shows boxplots gathered from two simulations in table 3. In the two plots shown, yet again the OLSAVS edges out the Lasso and has a much shorter variation in the box plot. Lasso with an exponential error does compete in the $n=100$ plot but then has a much larger gap when the sample size is doubled.
Table 4 uses uniformly distributed errors with zero mean. Once again, the same trend appears in the TRMSE values between the OLSAVS and Lasso estimates as before. The difference for the simulations remains mostly the same as for previous simulations for Lasso with an exponential distribution.
In figure 4, the variation for either method is large compared to figure 1, however, OLSAVS still has the smaller average TRMSE. Once the sample size increases, both the variation and the average TRMSE shrinks for each method, but the OLSAVS still maintains the advantage in each.
Tables 5 and 6, compare the OLSAVS method with Relaxed Lasso regression with normal errors and with exponential errors respectively. Much like the results for Lasso regression, The Relaxed Lasso appears to have a slight advantage when $k=1$. However, once $k$ increases, OLSAVS begins to have much smaller TRMSE values. However, the Relaxed Lasso is more competitive for normally distributed errors than the Lasso. The differences in table 6 show the OLSAVS method performing well in terms of bias of the regression coefficients. Additionally, the OLSAVS provides more consistent differences overall.
Figure 5 shows simulation results between OLSAVS and Relaxed Lasso with error type 1 while figure 6 shows results

Table 2. TRMSE and difference values for OLSAVS vs. Lasso for $e_{i} \sim t_{3}$

|  |  |  |  | TRMSE |  | Diff |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| n | p | k | $\psi$ | OLSAVS | Lasso | OLSAVS | Lasso |
| 100 | 20 | 1 | 0 | 1.3677 | 1.3362 | 0.0260 | 0.1882 |
| 100 | 20 | 1 | 0.3 | 1.3551 | 1.3317 | 0.1497 | 0.2031 |
| 100 | 20 | 1 | 0.9 | 1.3178 | 1.3169 | 0.6539 | 0.7482 |
| 100 | 20 | 19 | 0 | 1.4441 | 1.4442 | 0.0047 | 0.0130 |
| 100 | 20 | 19 | 0.3 | 1.4441 | 1.4590 | 0.0061 | 0.0142 |
| 100 | 20 | 19 | 0.9 | 1.4487 | 2.4245 | 1.8395 | 7.2843 |
| 100 | 50 | 1 | 0.3 | 1.3704 | 1.3311 | 0.2075 | 0.2549 |
| 100 | 50 | 1 | 0.9 | 1.3060 | 1.3015 | 0.9523 | 0.6991 |
| 100 | 50 | 49 | 0 | 1.8360 | 1.8418 | 0.0076 | 0.0076 |
| 100 | 50 | 49 | 0.3 | 1.9877 | 4.4636 | 0.0245 | 0.0245 |
| 100 | 50 | 49 | 0.9 | 2.2778 | 9.6242 | 8.8102 | 43.9438 |
| 200 | 40 | 1 | 0 | 1.3453 | 1.3165 | 0.0351 | 0.1550 |
| 200 | 40 | 1 | 0.3 | 1.3324 | 1.3163 | 0.1351 | 0.1707 |
| 200 | 40 | 1 | 0.9 | 1.3075 | 1.3033 | 0.8243 | 0.5671 |
| 200 | 40 | 39 | 0 | 1.4351 | 1.4355 | 0.0036 | 0.0107 |
| 200 | 40 | 39 | 0.3 | 1.4351 | 2.5826 | 0.0048 | 0.0376 |
| 200 | 40 | 39 | 0.9 | 1.8981 | 6.7108 | 5.7510 | 32.8242 |
| 200 | 100 | 1 | 0 | 1.3837 | 1.3189 | 0.0460 | 0.1847 |
| 200 | 100 | 1 | 0.3 | 1.3559 | 1.3153 | 0.1774 | 0.2115 |
| 200 | 100 | 1 | 0.9 | 1.3096 | 1.2998 | 1.1920 | 0.2210 |
| 200 | 100 | 99 | 0 | 1.8397 | 1.8487 | 0.0058 | 0.0147 |
| 200 | 100 | 99 | 0.3 | 3.3891 | 11.4710 | 0.1501 | 0.4976 |
| 200 | 100 | 99 | 0.9 | 3.8424 | 27.4098 | 17.2990 | 93.9071 |



Figure 2. Box plots for table 2 showing simulation results for OLSAVS vs. Lasso regression with error type 2


Figure 3. Box plots for table 3 showing simulation results for OLSAVS vs. Lasso regression with error type 3

Table 3. TRMSE and difference values for OLSAVS vs. Lasso for $e_{i} \sim \exp (1)-1$

|  |  |  |  | TRMSE |  | Diff |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| n | p | k | $\psi$ | OLSAVS | Lasso | OLSAVS | Lasso |
| 100 | 20 | 1 | 0 | 1.0589 | 1.0366 | 0.0208 | 0.1416 |
| 100 | 20 | 1 | 0.3 | 1.0287 | 1.0099 | 0.1123 | 0.1533 |
| 100 | 20 | 1 | 0.9 | 0.9913 | 0.9900 | 0.6264 | 0.6519 |
| 100 | 20 | 19 | 0 | 1.1055 | 1.1063 | 0.0033 | 0.0118 |
| 100 | 20 | 19 | 0.3 | 1.1055 | 1.1583 | 0.0046 | 0.0148 |
| 100 | 20 | 19 | 0.9 | 1.1335 | 2.3835 | 1.5415 | 7.5833 |
| 100 | 50 | 1 | 0 | 1.0948 | 1.0456 | 0.0232 | 0.1804 |
| 100 | 50 | 1 | 0.3 | 1.0767 | 1.0414 | 0.1689 | 0.2039 |
| 100 | 50 | 1 | 0.9 | 1.0263 | 1.0188 | 0.9343 | 0.5814 |
| 100 | 50 | 49 | 0 | 1.4078 | 1.4105 | 0.0045 | 0.0147 |
| 100 | 50 | 49 | 0.3 | 1.5744 | 4.3752 | 0.0153 | 0.0520 |
| 100 | 50 | 49 | 0.9 | 2.1137 | 9.5345 | 8.8138 | 44.0251 |
| 200 | 40 | 1 | 0 | 1.0298 | 1.0071 | 0.0202 | 0.1201 |
| 200 | 40 | 1 | 0.3 | 1.0202 | 1.0037 | 0.1063 | 0.1331 |
| 200 | 40 | 1 | 0.9 | 0.9969 | 0.9925 | 0.7664 | 0.4067 |
| 200 | 40 | 39 | 0 | 1.1034 | 1.1048 | 0.0026 | 0.0097 |
| 200 | 40 | 39 | 0.3 | 1.1034 | 2.4771 | 0.0035 | 0.0380 |
| 200 | 40 | 39 | 0.9 | 1.7076 | 6.7040 | 5.8761 | 32.8648 |
| 200 | 100 | 1 | 0 | 1.0522 | 1.0122 | 0.0371 | 0.0371 |
| 200 | 100 | 1 | 0.3 | 1.0356 | 1.0101 | 0.1402 | 0.1657 |
| 200 | 100 | 1 | 0.9 | 1.0067 | 0.9953 | 0.9636 | 0.0896 |
| 200 | 100 | 99 | 0 | 1.4138 | 1.4192 | 0.0037 | 0.0131 |
| 200 | 100 | 99 | 0.3 | 3.1658 | 11.2754 | 0.1451 | 0.4836 |
| 200 | 100 | 99 | 0.9 | 3.6705 | 27.4919 | 17.5606 | 93.9063 |

Table 4. TRMSE and difference values for OLSAVS vs. Lasso for $e_{i} \sim \operatorname{uniform}(-1,1)$

|  |  |  |  | TRMSE |  | Diff |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| n | p | k | $\psi$ | OLSAVS | Lasso | OLSAVS | Lasso |
| 100 | 20 | 1 | 0 | 0.6087 | 0.5952 | 0.0112 | 0.0843 |
| 100 | 20 | 1 | 0.3 | 0.6024 | 0.5945 | 0.0662 | 0.0900 |
| 100 | 20 | 1 | 0.9 | 0.5905 | 0.5879 | 0.5773 | 0.4067 |
| 100 | 20 | 19 | 0 | 0.6360 | 0.6369 | 0.0023 | 0.0103 |
| 100 | 20 | 19 | 0.3 | 0.6360 | 0.9626 | 0.0031 | 0.0308 |
| 100 | 20 | 19 | 0.9 | 0.7599 | 2.3075 | 1.3762 | 7.8005 |
| 100 | 50 | 1 | 0 | 0.6265 | 0.5999 | 0.0177 | 0.1043 |
| 100 | 50 | 1 | 0.3 | 0.6169 | 0.5955 | 0.0943 | 0.1160 |
| 100 | 50 | 1 | 0.9 | 0.5917 | 0.5826 | 0.7318 | 0.2120 |
| 100 | 50 | 49 | 0 | 0.8221 | 0.8299 | 0.0028 | 0.0127 |
| 100 | 50 | 49 | 0.3 | 1.0467 | 4.3592 | 0.0121 | 0.0573 |
| 100 | 50 | 49 | 0.9 | 1.9089 | 9.4600 | 9.0251 | 44.0993 |
| 200 | 40 | 1 | 0 | 0.6078 | 0.5907 | 0.0127 | 0.0697 |
| 200 | 40 | 1 | 0.3 | 0.6016 | 0.5904 | 0.0604 | 0.0760 |
| 200 | 40 | 1 | 0.9 | 0.5886 | 0.5836 | 0.4935 | 0.1028 |
| 200 | 40 | 39 | 0 | 0.6506 | 0.6506 | 0.0013 | 0.0088 |
| 200 | 40 | 39 | 0.3 | 0.6506 | 2.5639 | 0.0018 | 0.0388 |
| 200 | 40 | 39 | 0.9 | 1.4729 | 6.8174 | 5.9630 | 33.0055 |
| 200 | 100 | 1 | 0 | 0.6130 | 0.5903 | 0.0123 | 0.0831 |
| 200 | 100 | 1 | 0.3 | 0.5967 | 0.5967 | 0.0898 | 0.0973 |
| 200 | 100 | 1 | 0.9 | 0.5902 | 0.5830 | 0.5646 | 0.0216 |
| 200 | 100 | 99 | 0 | 0.8125 | 0.8243 | 0.0021 | 0.0119 |
| 200 | 100 | 99 | 0.3 | 2.9615 | 11.3001 | 0.1460 | 0.4744 |
| 200 | 100 | 99 | 0.9 | 3.5458 | 27.3187 | 17.4181 | 93.9087 |



Figure 4. Box plots for table 4 showing simulation results for OLSAVS vs. Lasso regression with error type 4


Figure 5. Box plots for table 5 showing simulation results for OLSAVS vs. Relaxed Lasso regression with error type 1

Table 5. TRMSE and difference values for OLSAVS vs. Relax Lasso for $e_{i} \sim N(0,1)$

|  |  |  |  | TRMSE |  | Diff |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| n | p | k | $\psi$ | OLSAVS | R Lasso | OLSAVS | R Lasso |
| 100 | 20 | 1 | 0 | 1.05004 | 1.03955 | 0.01180 | 0.04069 |
| 100 | 20 | 1 | 0.3 | 1.04307 | 1.03993 | 0.11183 | 0.11699 |
| 100 | 20 | 1 | 0.9 | 1.00809 | 1.00586 | 0.63942 | 0.67805 |
| 100 | 20 | 19 | 0 | 1.11518 | 1.11527 | 0.00315 | 0.00298 |
| 100 | 20 | 19 | 0.3 | 1.11518 | 1.12323 | 0.00384 | 0.00384 |
| 100 | 20 | 19 | 0.9 | 1.14518 | 1.44021 | 1.54564 | 7.55956 |
| 100 | 50 | 1 | 0 | 1.10563 | 1.08454 | 0.02597 | 0.05857 |
| 100 | 50 | 1 | 0.3 | 1.09096 | 1.08971 | 0.16601 | 0.16668 |
| 100 | 50 | 1 | 0.9 | 1.03093 | 1.02483 | 0.95682 | 0.60006 |
| 100 | 50 | 49 | 0 | 1.42151 | 1.42127 | 0.00415 | 0.00446 |
| 100 | 50 | 49 | 0.3 | 1.57851 | 3.89696 | 0.01567 | 0.01567 |
| 100 | 50 | 49 | 0.9 | 2.12296 | 4.80282 | 8.87674 | 45.66343 |
| 200 | 40 | 1 | 0 | 1.03855 | 1.03304 | 0.02126 | 0.02466 |
| 200 | 40 | 1 | 0.3 | 1.03295 | 1.03285 | 0.10517 | 0.10524 |
| 200 | 40 | 1 | 0.9 | 1.01005 | 1.00655 | 0.78436 | 0.42916 |
| 200 | 40 | 39 | 0 | 1.12531 | 1.12531 | 0.00267 | 0.00267 |
| 200 | 40 | 39 | 0.3 | 1.12531 | 1.80026 | 0.00386 | 0.12240 |
| 200 | 40 | 39 | 0.9 | 1.70105 | 3.57716 | 5.88467 | 33.41152 |
| 200 | 100 | 1 | 0 | 1.06191 | 1.05162 | 0.02599 | 0.03373 |
| 200 | 100 | 1 | 0.3 | 1.04883 | 1.04896 | 0.13610 | 0.13720 |
| 200 | 100 | 1 | 0.9 | 1.01643 | 1.00291 | 0.96788 | 0.96788 |
| 200 | 100 | 99 | 0 | 1.40641 | 1.40518 | 1.40518 | 0.00281 |
| 200 | 100 | 99 | 0.3 | 3.15561 | 10.08956 | 0.14074 | 0.92836 |
| 200 | 100 | 99 | 0.9 | 3.71833 | 9.86930 | 17.35998 | 97.95932 |

Note - R Lasso: Relaxed Lasso

Table 6. TRMSE and difference values for OLSAVS vs. Relaxed Lasso for $e_{i} \sim E X P(1)-1$

|  |  |  |  | TRMSE |  | Diff |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| n | p | k | $\psi$ | OLSAVS | R Lasso | OLSAVS | R Lasso |
| 100 | 20 | 1 | 0 | 1.36487 | 1.34437 | 0.02908 | 0.08346 |
| 100 | 20 | 1 | 0.3 | 1.34847 | 1.34209 | 1.34209 | 0.15530 |
| 100 | 20 | 1 | 0.9 | 1.30332 | 1.30211 | 0.62132 | 0.74412 |
| 100 | 20 | 19 | 0 | 1.44316 | 1.44326 | 0.00553 | 0.00605 |
| 100 | 20 | 19 | 0.3 | 1.44316 | 1.44720 | 0.00764 | 0.03972 |
| 100 | 20 | 19 | 0.9 | 1.43871 | 1.65548 | 1.83681 | 7.26831 |
|  | 50 | 1 | 0 | 1.44446 | 1.40636 | 0.03604 | 0.10807 |
| 100 | 50 | 1 | 0.3 | 1.41096 | 1.40477 | 0.20648 | 0.21225 |
| 100 | 50 | 1 | 0.9 | 1.34176 | 1.33864 | 0.95980 | 0.71164 |
| 100 | 50 | 49 | 0 | 1.85165 | 1.85164 | 0.00554 | 0.00698 |
| 100 | 50 | 49 | 0.3 | 2.00740 | 4.14320 | 4.14320 | 4.14320 |
| 100 | 50 | 49 | 0.9 | 2.31574 | 4.91246 | 8.71224 | 45.59930 |
| 200 | 40 | 1 | 0 | 1.33118 | 1.31585 | 0.03086 | 0.05388 |
| 200 | 40 | 1 | 0.3 | 1.31769 | 1.31472 | 0.13591 | 0.13865 |
| 200 | 40 | 1 | 0.9 | 1.29278 | 1.28860 | 0.84014 | 0.58889 |
| 200 | 40 | 39 | 0 | 1.44431 | 1.44436 | 0.00272 | 0.00272 |
| 200 | 40 | 39 | 0.3 | 1.44431 | 2.02409 | 0.00388 | 0.13289 |
| 200 | 40 | 39 | 0.9 | 1.91426 | 3.72937 | 5.77462 | 33.17816 |
| 200 | 100 | 1 | 0 | 1.35272 | 1.32707 | 0.04958 | 0.07102 |
| 200 | 100 | 1 | 0.3 | 1.33008 | 1.32910 | 0.18092 | 0.18116 |
| 200 | 100 | 1 | 0.9 | 1.28972 | 1.27810 | 1.19952 | 0.21588 |
| 200 | 100 | 99 | 0 | 1.80251 | 1.80167 | 0.00513 | 0.00464 |
| 200 | 100 | 99 | 0.3 | 3.31817 | 10.32727 | 0.14145 | 0.92488 |
| 200 | 100 | 99 | 0.9 | 3.76114 | 9.86319 | 17.36192 | 97.95763 |

Note - R Lasso: Relaxed Lasso


Figure 6. Box plots for table 6 showing simulation results for OLSAVS vs. Relaxed Lasso regression with error type 2.
with error type 2. Figure 6 follows closely to the results from figure 5. The Relaxed Lasso seems to have a much larger variation for error types 1 and 2 compared to OLASVS.

## 6. Real Data Example

Wisconsin nursing home data set provided by the Wisconsin Department of Health and Family Services (DHFS) was used as the real data example, see (Rosenberg et al., 2007). The goal of this data set is to utilize nursing home capacity. The years 2000 and 2001 were considered, with 362 and 355 facilities respectively. However, 10 observations were removed for containing missing values. The data set contains 12 variables, with total patient-years (TPY) being the response variable.

To determine how OLSAVS performs, the data set was split into testing and training sets. $60 \%$ of the data was allocated to be in the training set and $40 \%$ was in the testing set. The TRMSE was recorded and compared the values of the OLSAVS method to those of Lasso, Relaxed Lasso, and Elastic Net.
Table 7, summarizes the results. Each of these side-by-side results shows that the OLSAVS method edges out each of the other methods considered for this real-world example.

Table 7. TRMSE comparison for OLSAVS vs. common methods using Wisconsin nursing home data

| Method | OLSAVS | Elastic Net | OLSAVS | Lasso | OLSAVS | Relaxed Lasso |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| TRMSE | 7.956877 | 8.071371 | 8.007678 | 8.044224 | 8.007678 | 8.013652 |

## 7. The $R$ Package

The $R$ function used for the simulation, a function we used to produce graphs, and a function for performing the OLSAVS regression can be found under https://hasthika.github.io/olsvspack.txt. To load this package: use source("https://hasthika.github.io/olsvspack.txt")

## 8. Conclusions

The new method for variable selection, OLSAVS involves applying ordinary least squares to a subset of predictors selected from a specific variable selection method such as Lasso, Relax Lasso, or Elastic Net. We expected the OLSAVS to reduce the bias in the regression coefficients introduced by the shrinkage method and lead to the model being close-fitting while keeping the consistency of the reduced variance from the shrinkage method. Simulation results show that the OLSAVS method not only reduced the bias of the regression coefficient but also further reduced the variance of the estimates.

Furthermore, the OLSAVS method performs well in terms of prediction error as well. As discussed in section 5.2, the test root mean square errors when using OLSAVS for all error types studied are either significantly low or equal to the competing shrinkage method. It is interesting to notice that the prediction accuracy drastically decreases as the correlation between the predictors increases when using commonly used shrinkage methods. Prediction accuracy decreases further with the number of non-trivial predictors. OLSAVS method outperformed the other shrinkage method studied in both of the scenarios mentioned above and produced much lower test root mean square error values.

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